

A. Giovanidis 16.07.2014

Stochastic Geometry
modeling and analysis of wireless networks
(part I - theory)

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- ▶ [Baddeley \(2007\)](#), *Spatial Point Processes and their Applications*, Lecture Notes in Mathematics: Stochastic Geometry, Springer
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Outline for hour 1

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What is a Point Process.

Intro and Fundamentals

The Poisson Point Process

Operations preserving the law of Poisson

Conditioning a Point Process.

Palm Theory

Palm Measure, Slivnyak's Theorem, and Integration

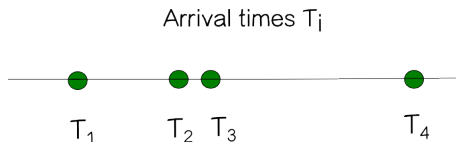
Marked Processes

A.1 - Intro and Fundamentals

Point Process in 1D (1)

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- ▶ A point process in one dimension ("time") is a useful model for the sequence of random times when a particular event occurs.
- ▶ Example "emergency calls in a hospital". Each call happens at an instant (point) of time.
- ▶ Given a period of time there will be a random number of such calls.



Point Process in 1D (2)

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- ▶ 1D Point Processes (PP) has a **natural ordering**.
- ▶ Arrival times: $T_1 < T_2 < \dots$ (dependent)
- ▶ Inter-arrival times: $S_i = T_{i+1} - T_i$ (independent for some)
- ▶ Cumulative counting process: $N_t = \sum_{i=1}^{\infty} \mathbf{1}_{\{T_i \leq t\}}$
- ▶ Interval counts: $N(a, b] = N_b - N_a$

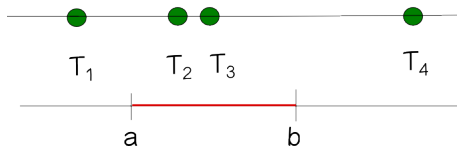
Point Process in 1D (2)

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- ▶ Interval counts: $N(a, b] = N_b - N_a$

N_t = total number of points up to time t , for all $t \geq 0$

Interval Counts $N(a, b] = 2$



Point Process in $\geq 2D$ (1)

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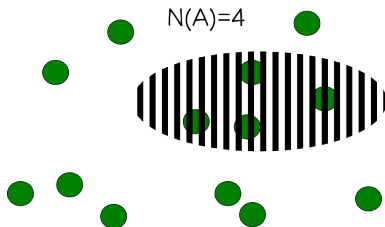
- ▶ In higher dimensions there is no natural ordering of points.
- ▶ No natural analogue for inter-arrival times S_i , counting process N_t .
- ▶ BUT! Generalise the **interval counts** to **region counts**.
- ▶ Superposition of PPs: $N(A) = N_1(A) + N_2(A)$ (additive).

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- ▶ Superposition of PPs: $N(A) = N_1(A) + N_2(A)$ (additive).

$N(A) = \text{number of points falling in } A, \quad A \subset \mathbb{R}^d$



Point Process in $\geq 2D$ (2)

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- ▶ It is interesting to study a PP using only the **vacancy indicators**.

$$\begin{aligned}V(A) &= \mathbf{1}\{N(A) = 0\} \\ &= \mathbf{1}\{\textit{there are no points falling in } A\}.\end{aligned}$$

- ▶ Superposition of PPs: $V(A) = V_1(A) \cdot V_2(A)$ (multiplicative).

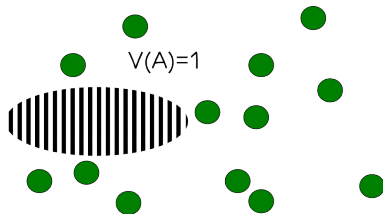
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Point Process Fundamentals (1)

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- ▶ A spatial **point process** Φ is a random, finite or countably-infinite collection of points in the space \mathbb{R}^d without accumulation points.
- ▶ Realization: Discrete subset $\phi = \{x_i\} \subset \mathbb{R}^d$. The indices of points follow some random enumeration.
- ▶ **Counting measure**: $N(A)$ gives the number of points of ϕ in A . $\mathbf{1}_{\{x \in A\}}$ is the Dirac measure at x .

$$N(A) = \sum_i \mathbf{1}_{\{x_i \in A\}}$$

Point Process Fundamentals (2)

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- ▶ For all real functions f defined on \mathbb{R}^d we have

$$\sum_i f(x_i) = \int_{\mathbb{R}^d} f(x) N(dx).$$

- ▶ No accumulation points $\Rightarrow N(A) < \infty$ for any bounded A .
- ▶ **Simple PP** $\Rightarrow \mathbb{P}[N(\{x\}) \leq 1, \forall x] = 1$, i.e. the points $\{x_i\}$ are pairwise different a.s..
- ▶ The distribution of a PP Φ is **entirely characterised** by the family of finite distributions $\{N(A_1), \dots, N(A_k)\}$, with A_1, \dots, A_k bounded.

Characteristics of the Point Process

- ▶ The intensity (mean) measure

$$\Lambda(A) = \mathbb{E}[N(A)].$$

- ▶ The empty (void) probability

$$\mathcal{V}(A) = \mathbb{P}[N(A) = 0], \quad A \in \mathcal{B}.$$

- ▶ The Laplace functional, for any function $f : \mathbb{R}^d \rightarrow \mathbb{R}_+$

$$\mathcal{L}(f) = \mathbb{E} \left[e^{-\int_{\mathbb{R}^d} f(x) N(dx)} \right].$$

It completely characterises the distribution of the PP. For $f(x) = \sum_{i=1}^k t_i \mathbf{1}_{(x \in A_i)} \Rightarrow \mathcal{L}(f) = \mathbb{E} \left[e^{-\sum_i t_i N(A_i)} \right]$ is the joint LT of the random vector $\{N(A_1), \dots, N(A_k)\}$.

Campbell's Formula

- Let f be a measurable function, then the equality holds

$$\mathbb{E} \left[\int_{\mathbb{R}^d} f(x) N(dx) \right] = \int_{\mathbb{R}^d} f(x) \Lambda(dx).$$

$$\stackrel{\text{(int. fun.)}}{=} \int_{\mathbb{R}^d} f(x) \beta(x) dx.$$

If the intensity measure Λ satisfies $\Lambda(A) = \int_A \beta(x) dx$ for some function β , then the latter is the **intensity function** of Φ .

$$\mathbb{P}[N(dx) > 0] \approx \mathbb{E}[N(dx)] \approx \beta(x) dx.$$

Stationarity and Isotropy

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A PP Φ is called:

1. **Stationary** (**homogeneous**) if the distribution of Φ is invariant with respect to **translation** of the origin, i.e.

$$\{X_i\} \stackrel{d.}{=} \{X_i - u\}, \forall u \in \mathbb{R}^d.$$

2. **Isotropic** if stationary & the distribution of Φ is invariant with respect to **rotation** of the origin.

A.2 - The Poisson Point Process

The Poisson Point Process

Definition:

The Poisson PP Φ of intensity measure Λ is defined by means of its finite-dimensional distributions:

$$\mathbb{P} \{N(A_1) = n_1, \dots, N(A_k) = n_k\} = \prod_{i=1}^k \left(e^{-\Lambda(A_i)} \frac{\Lambda(A_i)^{n_i}}{n_i!} \right),$$

for every $k = 1, 2, \dots$ and all **bounded, mutually disjoint sets** A_i , for $i = 1, \dots, k$.

If $\Lambda(dx) = \lambda dx \Rightarrow$ **homogeneous**.

Properties

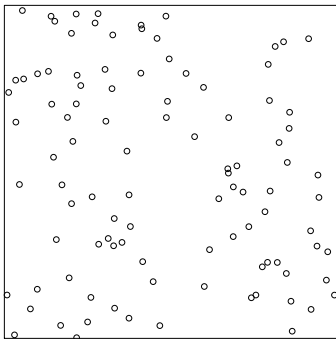
- ▶ For every compact set $A \subset \mathbb{R}^d$ the count $N(A)$ has a Poisson distribution with mean $\Lambda(A)$. Homogeneous case: $\lambda S(A)$.
- ▶ Complete Independence Property
If A_1, \dots, A_k are mutually disjoint compact sets, then $N(A_1), \dots, N(A_k)$ are independent random variables.
- ▶ Laplace Functional

$$\mathcal{L}(f) = e^{-\int_{\mathbb{R}^d} (1 - e^{-f(x)}) \Lambda(dx)}.$$

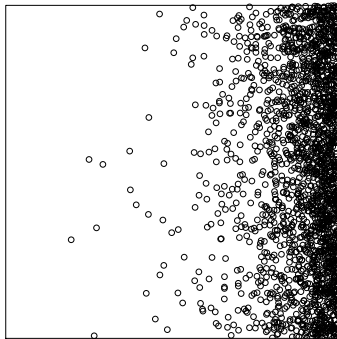
Examples

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Homogeneous Poisson Point Process



Non-homogeneous Poisson PP, $\beta(x,y)=\exp(10 \cdot x)$



Simulating the Poisson PP (1)

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Conditional Property

Consider a Poisson PP in \mathbb{R}^d with uniform intensity $\beta(x) = \lambda > 0$. Let $W \subset \mathbb{R}^d$ be any bounded region with $0 < \mathcal{S}(W) < \infty$.

Given $N(W) = n$, the conditional distribution of $N(B)$ for $B \subseteq W$ is **binomial**:

$$\mathbb{P}[N(B) = k | N(W) = n] = \binom{n}{k} p^k (1-p)^{n-k},$$

where $p = \frac{\mathcal{S}(B)}{\mathcal{S}(W)}$.

- ▶ Given there are n points of the Poisson PP in W , these are **conditionally independent** and **uniformly distributed** in W .

Simulating the Poisson PP (2)

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How to simulate a Poisson PP of intensity λ in W ?

1. First generate a random variable with Poisson distribution and mean $\lambda \mathcal{S}(W)$.
2. Given the realisation of the Poisson r.v. n , generate n independent uniform random point in W .

Simulating the Poisson PP (3)

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To see why this procedure produces a Poisson PP let us use the [Laplace functional](#)

$$\begin{aligned}
 \mathcal{L}(f) &= \mathbb{E} \left[\mathbf{1}_{\{N=0\}} + \mathbf{1}_{\{N>0\}} e^{-\sum_{k=1}^N f(X_i)} \right] \\
 &= e^{-\Lambda(W)} \sum_{k=0}^{\infty} \frac{\Lambda(W)^k}{k!} \left(\int_W e^{-f(x)} \frac{\Lambda(dx)}{\Lambda(W)} \right)^k \\
 &= e^{-\Lambda(W) + \int_W e^{-f(x)} \Lambda(dx)} \\
 &= e^{-\int_W (1 - e^{-f(x)}) \Lambda(dx)}.
 \end{aligned}$$

A.3 - Operations preserving the law of Poisson

Operation I: Mapping

Consider a fixed transformation $g : \mathbb{R}^d \rightarrow \mathbb{R}^d$ to each point x_i of Φ .

Let $g^{-1}(A)$ be bounded for each $A \in \mathbb{R}^d$ bounded.

- ▶ The resulting process has measure $N_M(A) = \sum_i \mathbf{1}_{\{g(x_i) \in A\}}$.
- ▶ For a Poisson PP with intensity Λ on \mathbb{R}^d , its transformation by g is also a Poisson PP with intensity $\Lambda(\cdot) = \Lambda(g^{-1}(\cdot))$.

Operation I: Mapping

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Examples

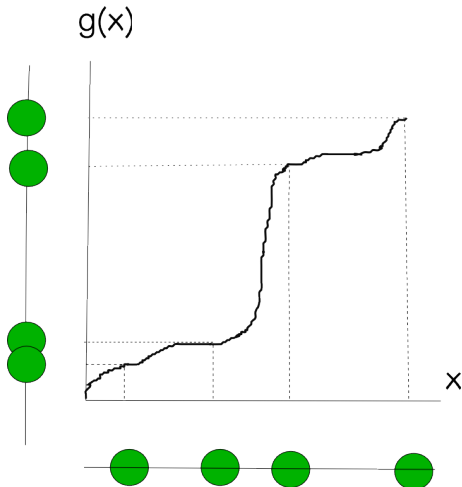
Dilation: $g(x) = \gamma x$, $\gamma > 0$, $\Lambda'(A) = \Lambda(A/\gamma)$.

Vector Translation: $g(x) = x + v$, $v \in \mathbb{R}^d$, $\Lambda'(A) = \Lambda(A - v) \stackrel{u.}{=} \lambda S(A)$.

Polar coordinates: $g(x) = (|x|, \angle x)$, $\Lambda'([0, r], [0, \theta]) = \lambda \pi r^2 \theta / (2\pi)$.

Mapping Example

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Operation II: Thinning

In **thinning** some fraction of points of the PP Φ are **deleted**.

- ▶ The thinning is a PP Φ_T given by $N_T(A) = \sum_i \Delta_i \mathbf{1}_{\{x_i \in A\}}$. The indicators Δ_i define which atom will stay.

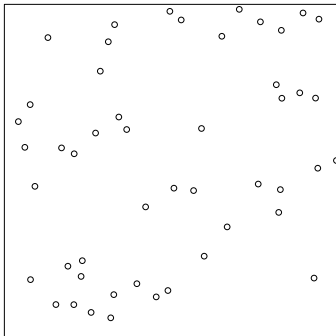
Independent Thinning

The r.v.'s Δ_i are **independent** and $\mathbb{P}[\Delta_i = 1] = \delta(x_i)$.

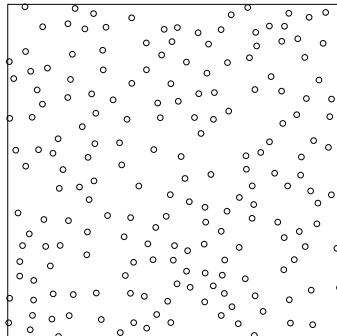
The independent thinning of a Poisson PP of intensity measure Λ with the **retention probability** δ yields a Poisson PP of intensity measure $\delta\Lambda$.

Dependent Thinning - Matérn Model I & II

rMaternI(500, 0.04)



rMaternII(500, 0.04)



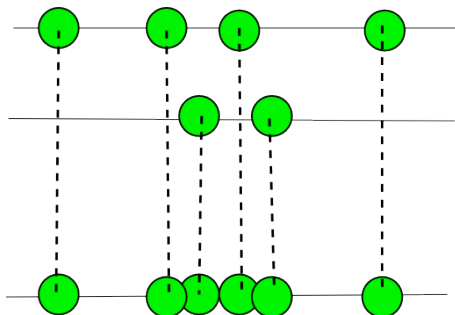
Operation III: Superposition

- ▶ The superposition of PPs Φ_k is defined as the sum $\Phi = \sum_k \Phi_k$.
- ▶ Viewed either as the **union of sets** or as the **union of measures**.
- ▶ If $N_{\Phi_1}(A)$ and $N_{\Phi_2}(A)$ are the number of points in region $A \subseteq \mathbb{R}^d$, then $N_{\Phi_1 + \Phi_2}(A) = N_{\Phi_1}(A) + N_{\Phi_2}(A)$.

The superposition of independent Poisson PPs with intensities Λ_k is a Poisson PP with intensity measure $\sum_k \Lambda_k$ **if and only if** the latter is a **locally finite measure**.

Superposition

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Operation IV: Cluster Formation

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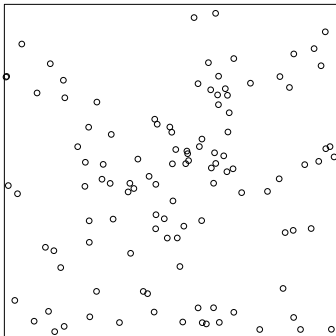
- ▶ Starting by a process Φ **replace** each point $x_i \in \Phi$ by a **random finite set of points** Z_i (the cluster).
- ▶ The superposition of all clusters yields $\Phi_C = \bigcup_i Z_i$.

Example: Matérn cluster process where the **parent** process is uniform Poisson and each cluster Z_i consists of a random Poisson number of points with mean μ , uniformly distributed in the disc $\mathcal{B}(x_i, \mu)$.

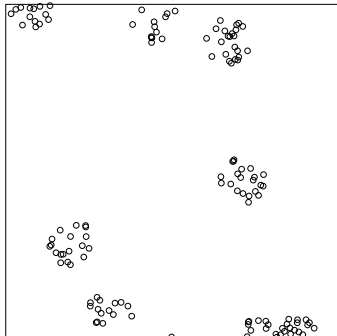
Matérn Cluster

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rMatClust(50, 0.07, 2)



rMatClust(5, 0.07, 20)



Outline for hour 2

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Palm Theory

Palm Measure, Slivnyak's Theorem, and Integration

Marked Processes

B.1 - Palm Theory

Motivation (1)

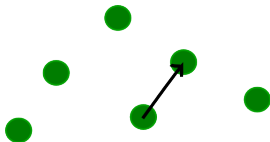
- ▶ Interested in properties relating to a **typical point** of the process.
- ▶ Conditional probabilities of events **given there is a point of the PP at a specific location**.
- ▶ Concept of **Palm distribution** and **Campbell-Mecke formula**.

Motivation (1)

- ▶ Interested in properties relating to a **typical point** of the process.
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- ▶ Concept of **Palm distribution** and **Campbell-Mecke formula**.

Q: What is the probability distribution of

- the **distance** from some atom of Φ to its **nearest neighbour atom**?



Nearest Neighbour Distance

Motivation (2)

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- ▶ Formally $R_x = \text{dist}(x, \Phi \setminus x)$, $x \in \Phi$. Find:

$$\mathbb{P}(R_x \leq r | x \in \Phi).$$

- ▶ **Non-elementary conditioning:** The event $\{x \in \Phi\}$ has 0 probability.

For the case of stationary PPs e.g. uniform Poisson

Condition on the event that there is a point of Φ at the origin 0.

Motivation (3)

- ▶ Let Φ be Poisson PP with intensity λ .
- ▶ Take a small neighbourhood around 0, e.g. the ball $\mathcal{B}(0, \epsilon)$.
- ▶ $R_\epsilon = \text{dist}(0, \Phi \setminus \mathcal{B})$
the distance from 0 to the nearest point of Φ outside \mathcal{B} .

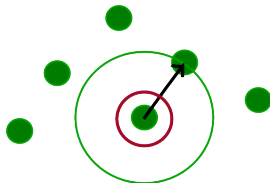
Motivation (3)

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- ▶ Let Φ be Poisson PP with intensity λ .
- ▶ Take a small neighbourhood around 0, e.g. the ball $\mathcal{B}(0, \epsilon)$.
- ▶ $R_\epsilon = \text{dist}(0, \Phi \setminus \mathcal{B})$
the distance from 0 to the nearest point of Φ **outside** \mathcal{B} .

$$\mathbb{P}[R_\epsilon > r | N(\mathcal{B}(0, \epsilon)) > 0] = \exp\{-\lambda\pi(r^2 - \epsilon^2)\}$$

$$\mathbb{P}[R_0 > r | 0 \in \Phi] = \exp\{-\lambda\pi r^2\}.$$



Nearest Neighbour Distance

B.2 - Palm Measure, Slivnyak's Theorem, and Integration

Palm Distribution

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C. Palm (1907-1951) for the study of telephone traffic.

- ▶ The **Palm probability** $\mathbb{P}^x(\Gamma)$ of an event Γ at location x is the conditional probability that the event Γ will occur, given $x \in \Phi$.
- ▶ If the events are restricted to the outcome events (counts of atoms) of a PP, we have the **Palm distribution** $\mathbf{P}^x(\Gamma)$ for the event Γ of a PP, defined as

$$\mathbf{P}^x(\Gamma) = \mathbb{P}^x(\Phi \in \Gamma).$$

Campbell Measure

- ▶ Reminder: Mean Radon measure

$$\Lambda(A) = \mathbb{E} \left[\int_{\mathbb{R}^d} \mathbf{1}_{\{x \in A\}} N(dx) \right] = \mathbb{E} [N(A)].$$

- ▶ We define the **Campbell measure**

$$C(A \times \Gamma) = \mathbb{E} \left[\int_{\mathbb{R}^d} \mathbf{1}_{\{x \in A\}} \mathbf{1}_{\{\phi \in \Gamma\}} N(dx) \right] = \mathbb{E} [N(A) \mathbf{1}_{\{\phi \in \Gamma\}}].$$

Palm Probability Measure

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Radon-Nikodým Theorem expresses the two measures in relation as

$$C(A \times \Gamma) = \int_A \mathbf{P}^x(\Gamma) \Lambda(dx).$$

$\mathbf{P}^x(\Gamma) = \frac{dC(\cdot \times \Gamma)}{d\Lambda(\cdot)}$ is the R-N derivative. It is also a probability measure.

Reduced Palm Measure

Let us remove the extra point x from the consideration. The **reduced Palm distribution** $\mathbf{P}^{!x}$ of a PP Φ is the distribution of $\Phi \setminus x$ under \mathbf{P}^x :

$$\mathbf{P}^{!x}(\Gamma) = \mathbf{P}^x(\Phi \setminus \{x\} \in \Gamma).$$

The relevant **reduced Campbell measure** of Φ is

$$\begin{aligned} C^!(A \times \Gamma) &= \mathbb{E} \left[\int_{\mathbb{R}^d} \mathbf{1}_{\{x \in A\}} \mathbf{1}_{\{\Phi \setminus x \in \Gamma\}} N(dx) \right] \\ &= \int_A \mathbf{P}^{!x}(\Gamma) \Lambda(dx) \end{aligned}$$

Slivnyak-Mecke Theorem (1)

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The **Poisson PP**, under the Palm distribution \mathbf{P}^x , behaves as if it were a Poisson PP **superimposed** with a fixed point at the location x

$$\mathbf{P}^x = \mathbf{P} * \Delta_x \quad (\text{convolution})$$

$$\Phi^x \stackrel{d.}{=} \Phi \cup \{x\}.$$

Slivnyak-Mecke Theorem (2)

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Theorem S-M: Let Φ be a PP with locally finite intensity measure Λ . Suppose the distributions \mathbf{P} and \mathbf{P}^x of Φ are related by

$$\mathbf{P}^x = \mathbf{P} * \Delta_x.$$

Then Φ is a **Poisson PP** with intensity measure Λ .

Theorem

Φ is a Poisson PP **if and only if**

$$\mathbf{P}^{!x} = \mathbf{P}.$$

Palm Distribution and Stationarity

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Palm distributions \mathbf{P}^x are equivalent under translation:

If Φ is a **stationary** point process in \mathbb{R}^d (e.g. Poisson), then

$$\Phi^x \stackrel{d.}{=} \Phi^0 + x,$$

where Φ^x again denotes a process governed by the Palm probability measure \mathbb{P}^x . More formally,

$$\mathbf{P}^x = T_x \mathbf{P}^0,$$

where T_x is the effect of translation by a vector x , $T_x N(A) = N(A - x)$.

Application: Nearest Neighbour Function

For a stationary PP Φ in \mathbb{R}^d , the nearest neighbour function G is the distribution of the distance from a typical point $x \in \Phi$ to the nearest other point of Φ

$$G(r) = \mathbb{P}^x (\text{dist}(x, \Phi \setminus \{x\}) \leq r) = \mathbf{P}^x (N(\mathcal{B}(x, r) \setminus \{x\}) > 0).$$

- ▶ Stationarity \Rightarrow the function does not depend on x .

Application: Nearest Neighbour Function

For a **stationary** PP Φ in \mathbb{R}^d , the **nearest neighbour function** G is the distribution of the distance from a **typical point** $x \in \Phi$ to the nearest other point of Φ

$$G(r) = \mathbb{P}^x (\text{dist}(x, \Phi \setminus \{x\}) \leq r) = \mathbf{P}^x (N(\mathcal{B}(x, r) \setminus \{x\}) > 0).$$

- ▶ Stationarity \Rightarrow the function does not depend on x .
- ▶ If **Poisson PP** we may use *Slivnyak's* theorem $\mathbf{P}^{!x} = \mathbf{P}$

$$\begin{aligned} G(r) &= \mathbb{P}^0 (\text{dist}(0, \Phi \setminus \{0\}) \leq r) = \mathbf{P} (\text{dist}(0, \Phi) \leq r). \\ &= 1 - \exp(-\lambda \pi r^2), \end{aligned}$$

equal to the **empty space function**.

Campbell-Mecke Formula

We want to calculate sums over Φ for functions that depend on the atom in x **and** on the entire PP.

For any $f(x, \Phi)$ that is integrable with respect to the Campbell measure,

$$\mathbb{E} \left[\sum_i f(x_i, \Phi) \right] = \int_{\mathbb{R}^d} \mathbb{E}^x [f(x, \Phi)] \Lambda(dx).$$

Campbell-Mecke Formula

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$$\mathbb{E} \left[\sum_i f(x_i, \Phi) \right] = \int_{\mathbb{R}^d} \mathbb{E}^x [f(x, \Phi)] \Lambda(dx).$$

For **Poisson PP**, use the *reduced Palm measure* and *Slivnyak's theorem*

$$\mathbb{E} \left[\sum_i f(x_i, \Phi \setminus \{x_i\}) \right] = \int_{\mathbb{R}^d} \mathbb{E} [f(x, \Phi)] \Lambda(dx).$$

C-M Application: Dependent Thinning A. Giovanidis 16.07.2014

Consider the Matern I PP: Results from thinning a Poisson PP

- ▶ Original Poisson PP Φ of intensity λ in \mathbb{R}^2 .
- ▶ Delete any $x \in \Phi$: $\text{dist}(x, \Phi \setminus \{x\}) \leq r_0$ (any point with neighbour in Φ closer than r_0). Resulting PP Ψ .
- ▶ $f(x, \Phi) = \mathbf{1}_{\{x \in A\}} \mathbf{1}_{\{\text{dist}(x, \Phi \setminus \{x\}) > r_0\}}$.
- ▶ $\mathbb{P}^x(\text{dist}(x, \Phi \setminus \{x\}) > r_0) = \exp(-\lambda\pi r_0^2)$.

C-M Application: Dependent Thinning A. Giovanidis 16.07.2014

Consider the Matern I PP: Results from thinning a Poisson PP

- ▶ Original Poisson PP Φ of intensity λ in \mathbb{R}^2 .
- ▶ Delete any $x \in \Phi$: $\text{dist}(x, \Phi \setminus \{x\}) \leq r_0$ (any point with neighbour in Φ closer than r_0). Resulting PP Ψ .
- ▶ $f(x, \Phi) = \mathbf{1}_{\{x \in A\}} \mathbf{1}_{\{\text{dist}(x, \Phi \setminus \{x\}) > r_0\}}$.
- ▶ $\mathbb{P}^x(\text{dist}(x, \Phi \setminus \{x\}) > r_0) = \exp(-\lambda\pi r_0^2)$.

$$\begin{aligned} \mathbb{E}[N(\Psi \cap A)] &= \mathbb{E}\left[\sum_i f(x_i, \Phi)\right] = \int_{\mathbb{R}^d} \mathbb{E}^x[f(x, \Phi)] \lambda dx = \\ &= \lambda \mathcal{S}(A) \exp\{-\lambda\pi r_0^2\} \Rightarrow \lambda_\Psi = \lambda \exp\{-\lambda\pi r_0^2\}. \end{aligned}$$

B.3 - Marked Processes

Marked PP (1)

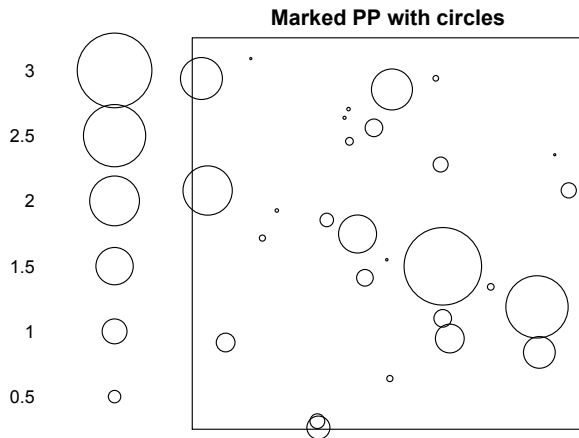
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- ▶ The points of a PP might be labelled with extra information called **marks**.
- ▶ e.g. If emergency calls are a PP on the line, each point might carry a label with the place of the call and the nature of emergency.
- ▶ Marked point is a pair: $\{x_i, m_i\}$, $m_i \in M \subseteq \mathbb{R}^k$.
- ▶ The marked PP is denoted by $\tilde{\Phi}$ and has counting measure $\tilde{N}(\cdot)$, i.e.

$$\tilde{N}(A \times M) = \sum_i \mathbf{1}_{\{x_i \in A, m_i \in M\}}.$$

Marked PP (2)

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Marked PP (3)

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- ▶ The corresponding projected PP of $\tilde{\Phi}$ on \mathbb{R}^d should be **locally finite**

$$\tilde{N}(A \times M) < \infty.$$

- ▶ A marked PP constructed by adding independent random marks to a Poisson PP is **equivalent to a Poisson PP in the product space**.
- ▶ **Thinning** a PP is formalised by a marked point process with marks in $\{0, 1\}$. $\{0\}$ =retain and $\{1\}$ =delete.

Marked Intensity

- ▶ The **intensity measure** of $\tilde{\Phi}$ is a measure on $A \times M$ determined by

$$\tilde{\Lambda}(A \times M) = \mathbb{E} \left[\tilde{N}(A \times M) \right] = \mathbb{E} \left[\sum_{(x,m) \in \tilde{\Phi}} \mathbf{1}_{\{x \in A\}} \mathbf{1}_{\{m \in M\}} \right].$$

- ▶ **Campbell's formula** takes the form

$$\mathbb{E} \left[\sum_{(x,m) \in \tilde{\Phi}} f(x, m) \right] = \int_{A \times M} f(x, m) \tilde{\Lambda}(dx, dm).$$

Independent Marking (1)

- ▶ **Independently Marked PP (i.m.PP)**: If given the locations of the points $\Phi = \{x_i\}$, the marks are mutually independent random vectors in M and if the conditional distribution satisfies

$$\mathbf{P}(m \in \cdot | \Phi) = \mathbf{P}(m \in \cdot | x) = Q_x(dm).$$

- ▶ The **Laplace functional** of an i.m.PP for all functions in \mathbb{R}^+ is

$$\begin{aligned}\mathcal{L}_{\tilde{\Phi}}(f) &= \mathbb{E} \left[e^{-\sum_i f(x_i, m_i)} \right] \\ &= e^{-\int_{\mathbb{R}^d} \left(1 - \int_M e^{-f(x, m)} Q_x(dm) \right) \Lambda(dx)}.\end{aligned}$$

Independent Marking (2)

The **reduced Campbell-Mecke formula** for i.m.PP gives

$$\mathbb{E} \left[\sum_i f(x_i, m_i, \Phi \setminus \{x_i\}) \right] = \int_{\mathbb{R}^d} \int_M \mathbb{E}^{!x} \left[f(x, m, \tilde{\Phi}) \right] Q_x(dm) \Lambda(dx),$$

and for the **Poisson PP** case

$$\mathbb{E} \left[\sum_i f(x_i, m_i, \Phi \setminus \{x_i\}) \right] = \int_{\mathbb{R}^d} \int_M \mathbb{E} \left[f(x, m, \tilde{\Phi}) \right] Q_x(dm) \Lambda(dx).$$

Marked Stationarity

- ▶ A marked point process on \mathbb{R}^d with marks in M is **stationary** if its distribution is **invariant under shifts of \mathbb{R}^d only**, $\forall v \in \mathbb{R}^d$

$$(x, m) \rightarrow (x + v, m).$$

- ▶ The location of the point x is shifted but the mark stays unchanged.
- ▶ For a stationary marked PP, the intensity measure takes the form

$$\mathbb{E} \left[\tilde{N}(A \times M) \right] = \lambda \mathcal{S}(A) Q_0(M).$$

λ = expected number of points per unit area, $Q_0(M)$ = **distribution of the typical mark** (probability measure).

END OF PART I